

Confidence Regions for the Central Orientation of Random Rotations

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INTRODUCTION

Though point estimation for random rotations has received considerable attention (see [3] for a review), inferential techniques are more limited. Here we derive the limiting distribution for extrinsic L_p estimators, which can be used to form asymptotic confidence regions (CR). Additionally we propose a nonparametric bootstrap CR that achieves close to nominal coverage rates for small n . In a simulation study we compare our methods to already existing ones in the literature.

SO(3) DATA AND ITS USES

- Orientation data in three dimensions can be represented by 3×3 orthogonal matrices with determinant one; the group of all such matrices is the rotation group denoted $SO(3)$.
- Some areas of application: structural geology, kinematics, material sciences, computer vision
- We consider the location model in $SO(3)$ given by

$$\mathbf{R}_i = \mathbf{S}\mathbf{E}_i \quad (1)$$

where $\mathbf{E}_i \in SO(3)$ are i.i.d. directionally symmetric perturbations of the central orientation $\mathbf{S} \in SO(3)$.

- (1) is analogous to the model $y_i = \mu + e_i$ on \mathbb{R} for $e_i \stackrel{iid}{\sim} (0, \sigma^2)$

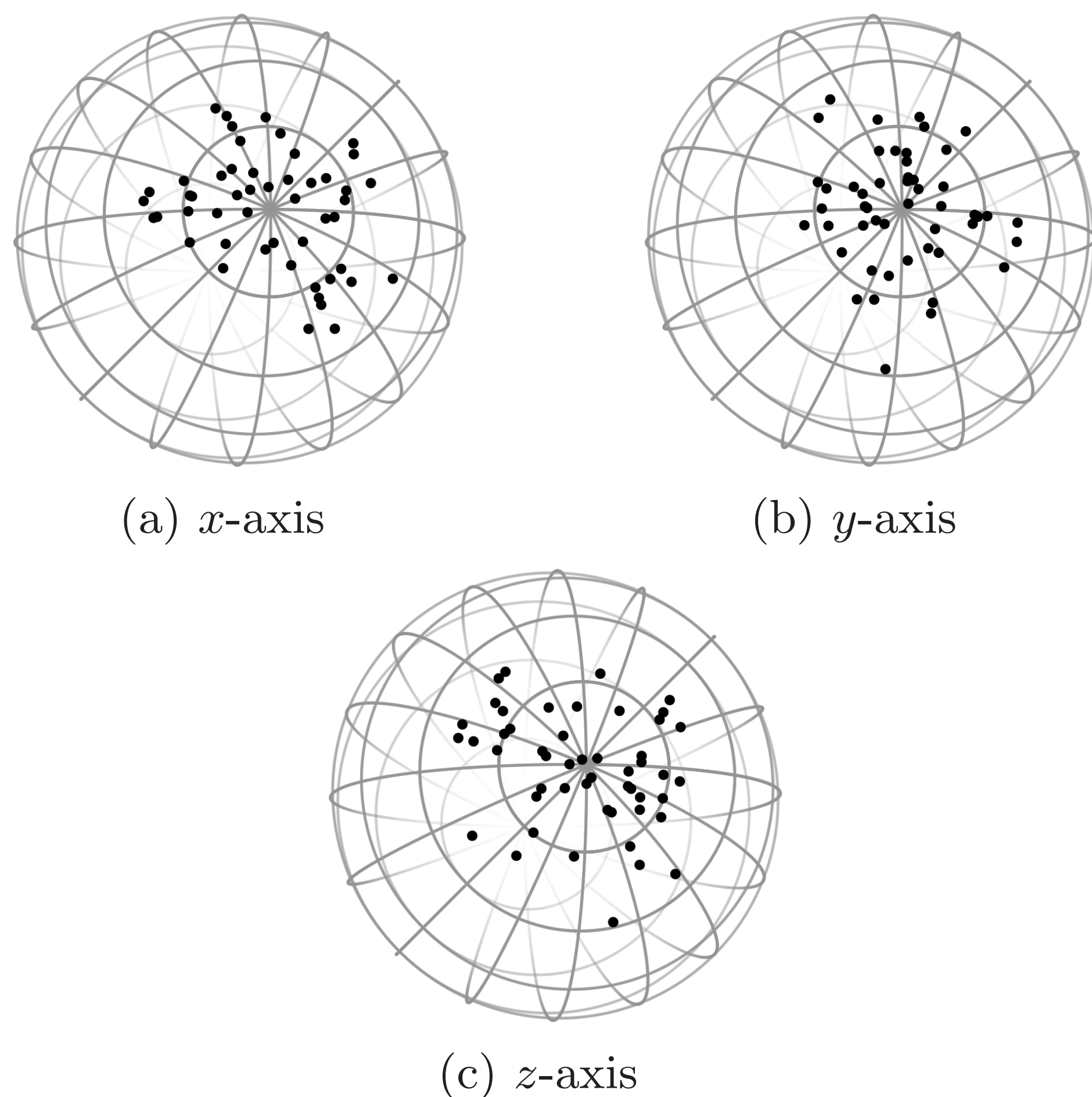


Figure 1: A random sample from the $SO(3)$ location model with $\mathbf{S} = \mathbf{I}_{3 \times 3}$ and $\mathbf{E}_i \stackrel{iid}{\sim} \text{Cayley}(\kappa = 50)$

- Two approaches to estimating \mathbf{S} in (1): the extrinsic approach based on the Euclidean distance and the intrinsic approach which makes use of innate $SO(3)$ topology by using Riemannian distance, see [3] for more details.

EXTRINSIC L_p ESTIMATORS

- Extrinsic (projected) estimators are based on the Euclidean distance metric defined for rotations $\mathbf{R}_1, \mathbf{R}_2 \in SO(3)$ as

$$d_E(\mathbf{R}_1, \mathbf{R}_2) = \|\mathbf{R}_1 - \mathbf{R}_2\|_F$$

where $\|\mathbf{A}\|_F^2 = \text{tr}(\mathbf{A}^\top \mathbf{A})$ is the Frobenius norm of the matrix \mathbf{A} and $\text{tr}(\cdot)$ is the matrix trace

- The projected median is defined as

$$\hat{\mathbf{S}}_1 = \arg \min_{\mathbf{S} \in SO(3)} \sum_{i=1}^n d_E(\mathbf{R}_i, \mathbf{S})$$

- The projected mean is defined as

$$\hat{\mathbf{S}}_2 = \arg \min_{\mathbf{S} \in SO(3)} \sum_{i=1}^n d_E^2(\mathbf{R}_i, \mathbf{S}) = \arg \max_{\mathbf{S} \in SO(3)} \text{tr}(\mathbf{S}^\top \bar{\mathbf{R}})$$

where $\bar{\mathbf{R}} = \sum_{i=1}^n \mathbf{R}_i / n$.

LARGE SAMPLE THEORY

Proposition: Assume $\mathbf{R}_1, \mathbf{R}_2, \dots, \mathbf{R}_n$ are a sample of i.i.d. observations from a rotationally symmetric location model with central orientation \mathbf{S} , $\hat{\mathbf{S}}_P = \arg \min_{\mathbf{S} \in SO(3)} \sum_{i=1}^n d_E^p(\mathbf{R}_i, \mathbf{S})$ is a consistent estimator for \mathbf{S} and $\hat{\mathbf{h}}_p$ satisfies $\exp[\Phi(\hat{\mathbf{h}}_p)] = \mathbf{S}^\top \hat{\mathbf{S}}_P$. Then

$$\sqrt{n} \hat{\mathbf{h}}_p \xrightarrow{d} MVN_3 \left(\mathbf{0}, \frac{c}{2d^2} \mathbf{I} \right)$$

as $n \rightarrow \infty$ or equivalently

$$\frac{2nd^2}{c} \|\hat{\mathbf{h}}_p\| \rightarrow \chi_3^2 \quad (2)$$

as $n \rightarrow \infty$. When $p = 1$, $\hat{\mathbf{S}}_P = \hat{\mathbf{S}}_1$,

$$c = \frac{1}{6} E[1 + \cos(r)] \quad \text{and} \quad d = \frac{1}{12} E \left[\frac{1 + 3 \cos(r)}{\sqrt{1 - \cos(r)}} \right];$$

assuming that $\hat{\mathbf{S}}_1 \neq \mathbf{R}_i$ for all i guarantees d is defined. When $p = 2$, $\hat{\mathbf{S}}_P = \hat{\mathbf{S}}_2$,

$$c = \frac{2}{3} E[1 - \cos^2(r)] \quad \text{and} \quad d = \frac{1}{3} E[1 + 2 \cos(r)].$$

From (2), the set of rotations \mathbf{R} that satisfy

$$\frac{2nd^2}{\hat{c}} d_R(\hat{\mathbf{S}}_P, \mathbf{R})^2 < \chi_{3,1-\alpha}^2$$

form a $100(1-\alpha)\%$ CR for the central orientation \mathbf{S} where \hat{c} and \hat{d} are consistent estimators of c and d , respectively.

SIMULATION STUDY

- The convergence rate of (2) is different for $\hat{\mathbf{S}}_1$ and $\hat{\mathbf{S}}_2$, see Figure 2 (right).
- Different c, d values for $\hat{\mathbf{S}}_1$ and $\hat{\mathbf{S}}_2$ result in different asymptotic variances, Figure 3 (below) illustrates the effect on CR coverage rates.

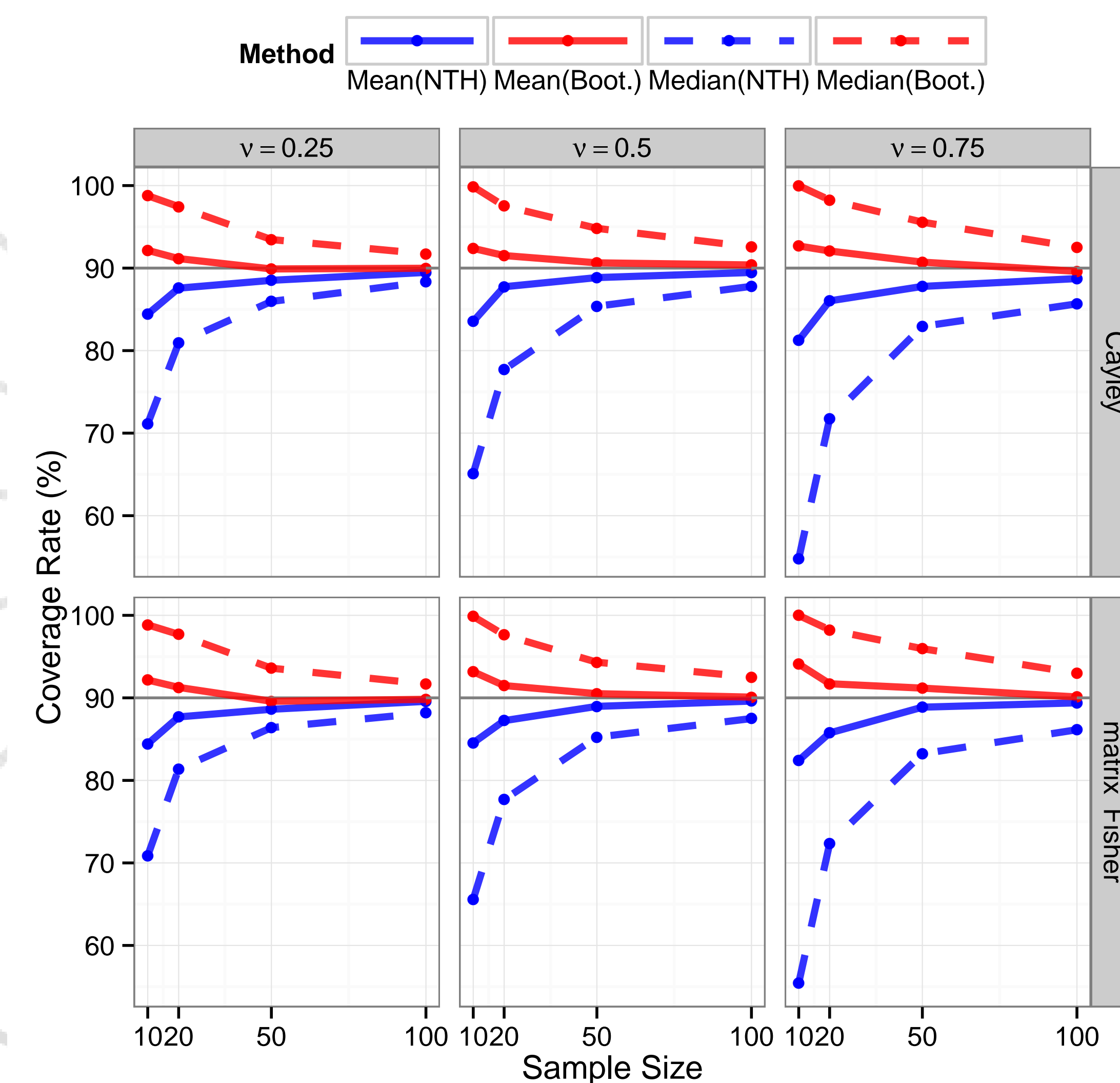


Figure 3: Coverage rates for 90% CRs using $\hat{\mathbf{S}}_1$ (dashed) and $\hat{\mathbf{S}}_2$ (solid)

- Existing CR methods based on $\hat{\mathbf{S}}_2$ rely on eigenvector asymptotics, our result is moment-based. No other CRs centered at $\hat{\mathbf{S}}_1$ exist.
- Figure 4 (right) compares existing methods (denoted “Eigen”) to our method (denoted “Moment”); for normal theory see [1], for bootstrap see [2].

DISCUSSION

- $\hat{\mathbf{S}}_1$ has a larger asymptotic variance than $\hat{\mathbf{S}}_2$ for rotationally symmetric samples.
- Critical value based on $\hat{\mathbf{S}}_1$ converges more slowly than for $\hat{\mathbf{S}}_2$.

FURTHER RESEARCH

- Explore estimator and CR behavior under more general assumptions
- Extend the result to intrinsic estimators
- Quantify differences in extrinsic versus intrinsic estimation choice

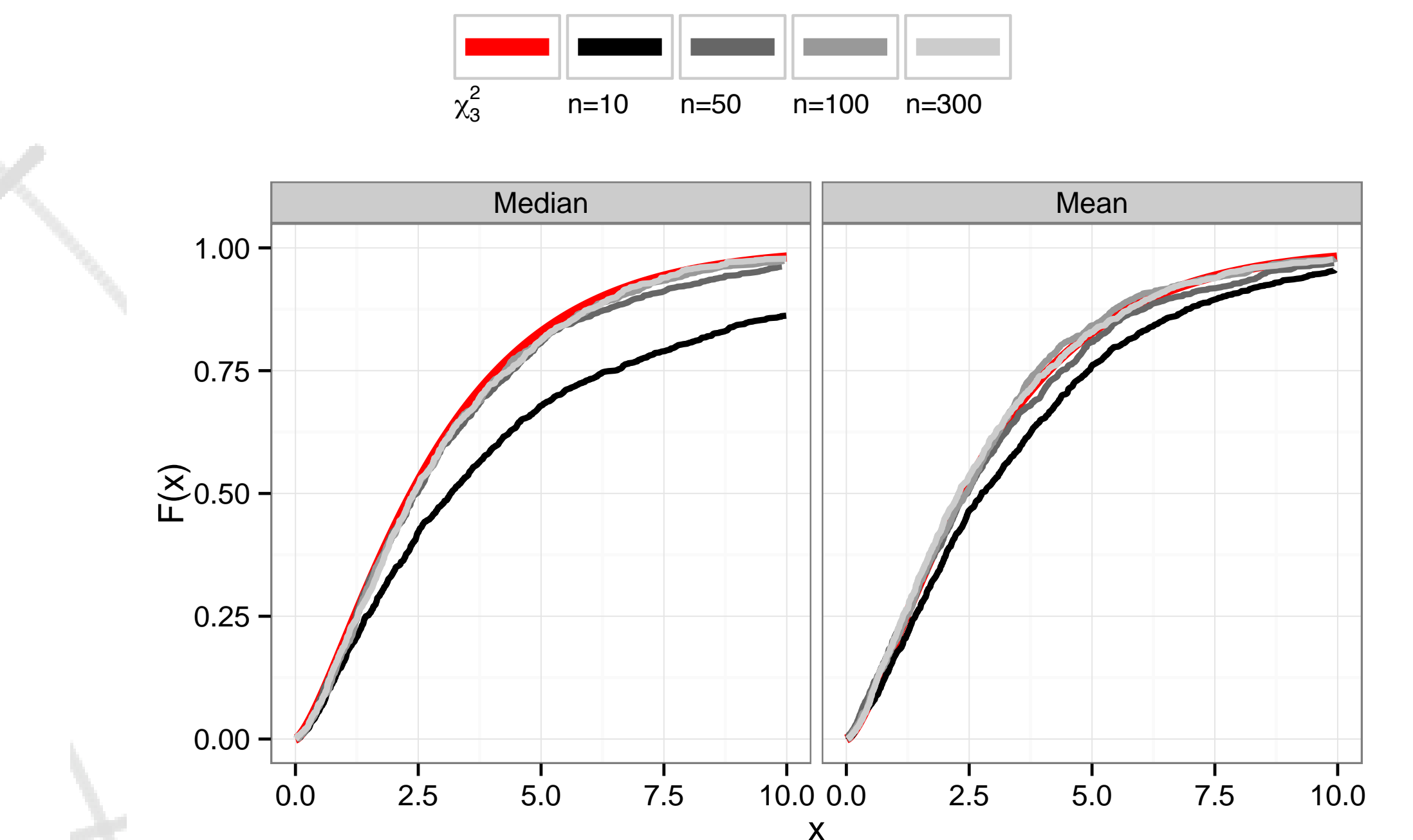


Figure 2: Critical value ECDF and theoretical limiting χ_3^2 distribution

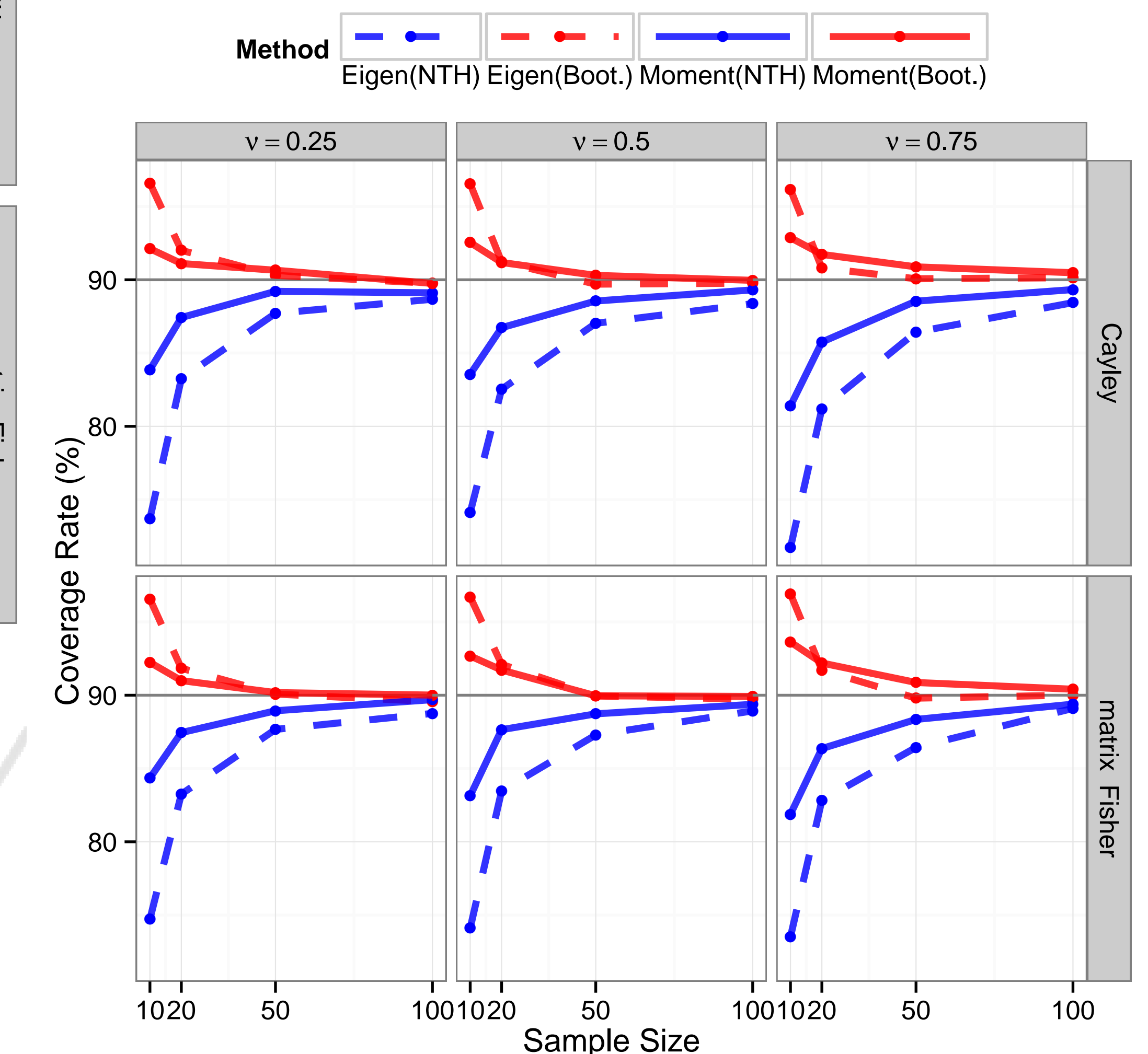


Figure 4: Coverage rates for 90% CRs using $\hat{\mathbf{S}}_2$ based on new (solid) and existing methods (dashed)

- All CR methods monotonically approach nominal coverage rate as n increases.
- Moment-based methods generally perform best, especially for small n .

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- [2] N. Fisher, P. Hall, B. Jing, and A. Wood. Improved pivotal methods for constructing confidence regions with directional data. *JASA*, 91(435):1062–1070, 1996.
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