Inference for the Central Direction of Random Rotations in SO(3)

Joint Statistical Meetings 2013

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OUTLINE

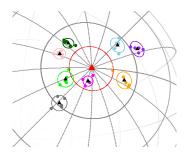
Introduction CoDA Rotation Data

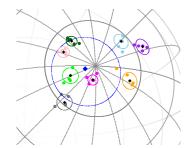
Confidence Regions
Setup
Result
Pivotal Bootstrap
Simulation Study
Extension

Future Work

CODA POSTER

- ► My CoDA poster included the following illustration of confidence regions for *SO*(3) data
- ► In this talk I provide the theoretical justification for these plots





▶ $R_1, ..., R_n \in SO(3)$ random sample from

$$\mathbf{R}_i = \mathbf{S}\mathbf{E}_i, \quad i = 1, \dots, n, \tag{1}$$

- Where
 - ► SO(3) collection of all 3×3 matrices R with det(R) = 1 and $R^{\top}R = I$
 - ▶ $S \in SO(3)$ parameter measuring central tendency
 - ► $E_1, ..., E_n \in SO(3)$ i.i.d. random rotations (symmetrically perturb S)
- ▶ Real line analog: $Y_i = \mu + e_i$ for $\mu \in \mathbb{R}$ and $e_i \in \mathbb{R}$ additive error symmetrically distributed around zero

R associated with skew-symmetric matrix $\Phi(W)$

$$\mathbf{\Phi}(\mathbf{W}) = \begin{bmatrix} 0 & -w_3 & w_2 \\ w_3 & 0 & -w_1 \\ -w_2 & w_1 & 0 \end{bmatrix}$$

for $W = (w_1, w_2, w_3) \in \mathbb{R}^3$ according to

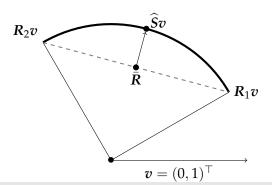
$$R = \exp[\Phi(\mathbf{W})] = \cos(r)\mathbf{I} + \sin(r)\Phi(\mathbf{U}) + (1 - \cos r)\mathbf{U}\mathbf{U}^{\top}$$
 (2)

- ► $r = \|W\|_F$ and $U = W/\|W\|_F$
- ► *r* distributed symmetrically about 0 independent of *U* distributed uniformly on unit sphere
- ► *r* and *U* termed misorientation angle and axis, respectively

► The projected mean is the most developed estimator

$$\widehat{S} = \operatorname*{arg\,min}_{S \in SO(3)} \sum_{i=1}^{n} d_E^2(R_i, S) = \operatorname*{arg\,max}_{S \in SO(3)} \operatorname{tr}(S^{\top} \overline{R})$$

 $ightharpoonup ar{R} = rac{1}{n} \sum_{i=1}^{n} R_i \text{ and } d_E(R_1, R_2) = \|R_1 - R_2\|_F$



SETUP

- ▶ $R_1, ..., R_n \in SO(3)$ random sample from (1) with central orientation S
- Use estimator $\hat{S} = \underset{S \in SO(3)}{\arg \min} \sum_{i=1}^{n} d_E^2(R_i, S)$
- $lacktriangledown \Phi(\hat{h})$ skew-symmetric matrix associated with $S^{\top}\widehat{S}$

$$\exp[\Phi(\hat{\mathbf{h}})] = \mathbf{S}^{\top} \widehat{\mathbf{S}}$$

▶ Define $r_{\hat{\pmb{h}}} = \|\hat{\pmb{h}}\|_F$ and $\pmb{U}_{\hat{\pmb{h}}} = \hat{\pmb{h}}/\|\hat{\pmb{h}}\|_F$

ASYMPTOTIC RESULT

Under this setup

$$\sqrt{n}\hat{\boldsymbol{h}} \stackrel{d}{\to} MVN_3\left(\boldsymbol{0}, \frac{c}{2d^2}I\right)$$
 (3)

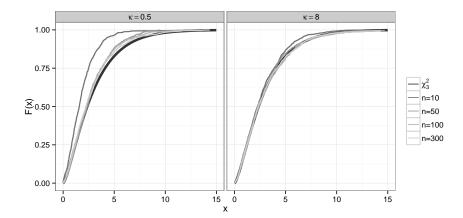
as $n \to \infty$ where

$$c = \frac{2}{3}E(1 - \cos^2 r)$$
 and $d = \frac{1}{3}E(1 + 2\cos r)$

or equivalently

$$\frac{2nd^2}{c}\|\hat{\boldsymbol{h}}\|_F^2 \stackrel{d}{\to} \chi_3^2$$

ECDF FOR PROJECTED MEAN



CONFIDENCE REGION

▶ By definition of the Riemannian distance d_R :

$$\|\hat{h}\|_F^2 = r_{\hat{h}}^2 = d_R(S, \widehat{S})^2$$

▶ A $100(1 - \alpha)$ % confidence region for *S*

$$\left\{ S \in SO(3) : \frac{2n\hat{d}^2}{\hat{c}} [d_R(S, \widehat{S})^2] < \chi^2_{3,1-\alpha} \right\}$$
 (4)

where

$$\hat{c} = \frac{1}{6n} \sum_{i=1}^{n} \left\{ 3 - \operatorname{tr} \left[\left(\widehat{S}^{\top} R_{i} \right)^{2} \right] \right\} \quad \text{and} \quad \hat{d} = \frac{1}{3n} \sum_{i=1}^{n} \operatorname{tr} \left(\widehat{S}^{\top} R_{i} \right)$$

• (4) describes confidence region centered at \hat{S} with radius

$$\sqrt{\frac{\hat{c}\chi_{3,1-\alpha}^2}{2n\hat{d}^2}}$$

PIVOTAL BOOTSTRAP¹

- 1. Randomly select n rotation matrices with replacement from the sample to form a bootstrap sample R_1^*, \ldots, R_n^*
- 2. Compute the projected mean of the bootstrap data set, \widehat{S}^* and form the test quantity $\frac{2n\widehat{d}^{*2}}{\widehat{c}^*}[d_R(\widehat{S},\widehat{S}^*)^2]$ where \widehat{c}^* and \widehat{d}^* are computed from the bootstrap sample by replacing R_i and \widehat{S} with R_i^* and \widehat{S}^*
- 3. Repeat steps 1 and 2, m times to obtain m values of the quantity $\frac{2n\hat{d}^{*2}}{\hat{c}^{*}}[d_{R}(\widehat{S},\widehat{S}^{*})^{2}]$
- 4. Define $\hat{q}_{1-\alpha}$ such that $P(\frac{2n\hat{d}^{*2}}{\hat{c}^*}[d_R(\widehat{S},\widehat{S}^*)^2] \leq q_{1-\alpha}^*) = 1-\alpha$
- 5. $\{S \in SO(3) : \frac{2n\hat{d}^2}{\hat{c}}[d_R(S,\widehat{S})^2] < \hat{q}_{1-\alpha}\}.$

 $^{^{1}\}mbox{First}$ proposed in Dr. Zhang's M.S. Thesis completed under Dr. Dan Nordman

OTHER CONFIDENCE REGION METHODS

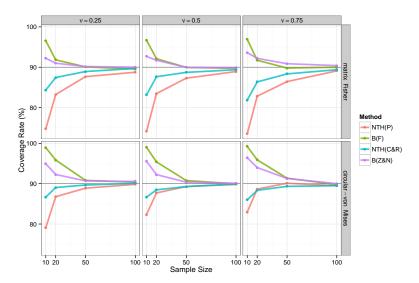
- ▶ Prentice (1986) and (1989) used asymptotics for eigenvalues and eigenvectors to construct a confidence region for \hat{S}
- ► Fisher, Hall, Jing and Wood (1996) use the Prentice statistic and a pivotal bootstrap procedure to achieve better coverage rates
- ► Chang and Rivest (2001) state a result more general than (3) that is difficult to implement

SIMULATION STUDY

Study parameters:

- ▶ Distributions: matrix Fisher, circular-von Mises
- ► Sample Sizes: n = 10, 20, 50, 100
- ► Circular Variances: $\nu = 0.25, 0.50$ and 0.75
- ► Simulated Samples: 10,000 per combination
- ▶ Bootstrap Sample Size: m = 300
- Error rate: $\alpha = 0.1$

COVERAGE RATE COMPARISON



PROJECTED MEDIAN

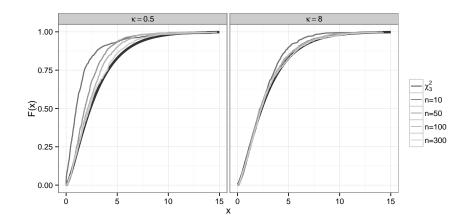
► The projected median

$$\widetilde{S} = \underset{S \in SO(3)}{\operatorname{arg min}} \sum_{i=1}^{n} d_{E}(R_{i}, S)$$

▶ (3) holds for \widetilde{S} with

$$c = \frac{1}{6}E(1 + \cos r) \quad \text{and} \quad d = \frac{1}{12}E\left(\frac{1 + 3\cos r}{\sqrt{1 - \cos r}}\right)$$

ECDF FOR PROJECTED MEDIAN



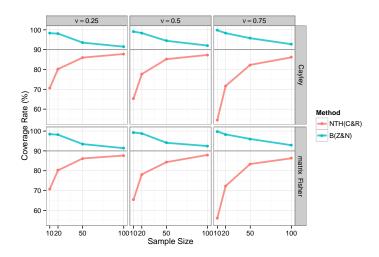
SIMULATION STUDY FOR MEDIAN

- $ightharpoonup \widetilde{S}$ cannot be written as a function of eigenvalues and eigenvectors
- ► The pivotal boostrap from before can still be used using by replacing \hat{c} and \hat{d} with

$$\tilde{c} = \frac{1}{12n} \sum_{i=1}^{n} \left[1 + \operatorname{tr}\left(\tilde{\boldsymbol{S}}^{\top} \boldsymbol{R}_{i}\right) \right] \quad \text{and} \quad \tilde{d} = \frac{\sqrt{2}}{24n} \sum_{i=1}^{n} \frac{3\operatorname{tr}(\tilde{\boldsymbol{S}}^{\top} \boldsymbol{R}_{i}) - 1}{\sqrt{3 - \operatorname{tr}(\tilde{\boldsymbol{S}}^{\top} \boldsymbol{R}_{i})}}$$

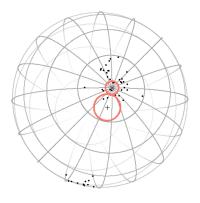
► Use same simulation parameters for the projected mean for the median

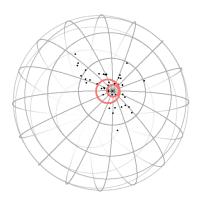
COVERAGE RATE SIMULATION \widetilde{S}



FUTURE WORK

► Look at more interesting cases, e.g. grain boundries





Thank you

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References:

- M. Prentice. Orientation statistics without parametric assumptions. *JRSS Series B*, 48(2):214-222, 1986.
- M. Prentice. Spherical regression on matched pairs of orientation statistics. *JRSS Series B*, 51(2):241-248, 1989.
- N. Fisher, P. Hall, B. Jing, and A. Wood. Improved pivotal methods for constructing confidence regions with directional data. *JASA*, 91(435):1062-1070, 1996.
- T. Chang and L. Rivest. M-estimation for location and regression in group models: A case study using stiefel manifold. *The Annals of Statistics*, 29(3): 784-814, 2001